#### **Generalized Linear Latent and Mixed Models**

# Sophia Rabe-Hesketh

University of California, Berkeley Cal

& Institute of Education, University of London



Australian Statistical Conference Adelaide, July 2012

. - p.1

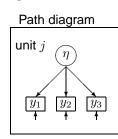
#### Basic idea of latent variable models

- What is a latent variable?
  - Random variable  $\eta_i$  for unit j whose realized values are hidden
  - (Hierarchical Bayes:  $\eta_i$  exchangeable, prior has free hyperparameters)
  - Properties must be indirectly inferred based on a statistical model connecting observed variables  $\mathbf{y}_j' = (y_{1j}, y_{2j}, \dots, y_{nj})$  to  $\eta_j$
- Basic construction principle of latent variable models:
   Conditional independence of observed variables given latent variable

$$\bullet \ \mathsf{Pr}(\mathbf{y}_j|\eta_j) \ = \ \prod_{i=1}^{n_j} \mathsf{Pr}(y_{ij}|\eta_j)$$

• Dependence among  $y_{ij}$  for unit j induced by  $\eta_i$ 

 $\Rightarrow$  Infer properties of  $\eta_j$  from  $Pr(y_j)$ 



#### **Outline**

- ► Classical latent variable models
  - Factor analysis
  - Item response theory
  - · Generalized linear mixed models
  - · Latent class models
  - (Skip structural equation models)
- ► Generalized linear latent and mixed models (GLLAMMs)
- Examples of GLLAMMs
  - · Several examples
  - Multilevel structural equation models

GLLAMM - p.2

#### Uses of latent variables

- Hypothetical constructs
- ► True variables (free from measurement error)
- Unobserved heterogeneity
- Missing data
- Device to induce dependence between different response types

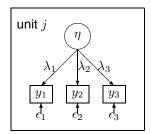
GLIAMM – p.3 GLIAMM – p.4

## Factor analysis

- $\blacktriangleright$  Common factor  $\eta_i$  underlies different continuous variables  $y_{ij}$ 
  - Spearman's (1904) factor g explains correlations between test scores in different subjects (French, English, Math, Music, etc.)
- ▶ Unidimensional factor model for *i*th variable for unit *j*:

$$y_{ij} = \beta_i + \lambda_i \eta_j + \epsilon_{ij}$$
 or  $E(y_{ij}|\eta_j) = \nu_{ij} = \beta_i + \lambda_i \eta_j$ 

- $\beta_i$  is intercept
- $\lambda_i$  is factor loading
- $\eta_i \sim N(0, \psi)$  is common factor
- $\epsilon_{ij} \sim N(0, \theta_i)$  is unique factor
- For identification, set  $\psi = 1$  or  $\lambda_1 = 1$



▶ Measurement model – It is often assumed that  $\lambda_i = 1$  and  $\theta_i = \theta$ 

GLLAMM - p.5

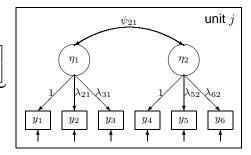
#### Multidimensional measurement model

- ► Several latent traits, e.g. verbal and spatial intelligence
- ▶ Model for vector of linear predictors  $\nu_i$

$$\nu_i = \beta + \Lambda \eta_i, \quad \eta_i \sim N(\mathbf{0}, \mathbf{\Psi})$$

▶ Ex: Independent clusters (between-item) two-dimensional model

$$\begin{bmatrix}
\nu_{1j} \\
\nu_{2j} \\
\nu_{3j} \\
\nu_{4j} \\
\nu_{5j} \\
\nu_{6j}
\end{bmatrix} = \begin{bmatrix}
\beta_1 \\
\beta_2 \\
\beta_3 \\
\beta_4 \\
\beta_5 \\
\beta_6
\end{bmatrix} + \begin{bmatrix}
1 & 0 \\
\lambda_{21} & 0 \\
\lambda_{31} & 0 \\
0 & 1 \\
0 & \lambda_{52} \\
0 & \lambda_{62}
\end{bmatrix} \underbrace{\begin{bmatrix}
\eta_{1j} \\
\eta_{2j} \\
\eta_{j}
\end{bmatrix}}_{\boldsymbol{\eta}_{j}}$$

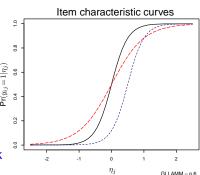


#### Item response model

- ▶ Measurement model for binary data
  - In educational testing, ability  $\eta_j$  explains performance in test items ( $y_{ij} = 1$ : correct,  $y_{ij} = 0$ : incorrect)
- ▶ Two-parameter logistic (2-PL) model for *i*th item for person *j*

$$logit[Pr(y_{ij}=1|\eta_j)] = \nu_{ij} = \beta_i + \lambda_i \eta_j$$

- $\beta_i$  is item intercept
- $\lambda_i$  is item discrimination
- $\eta_i \sim N(0, \psi)$  is person ability
- $-\beta_i/\lambda_i$  is item difficulty  $\Pr(y_{ij}=1|\eta_j=-\beta_i/\lambda_i)=0.5$
- Rasch model:  $\lambda_i = 1$
- ▶ Common factor model with logit link



# Generalized linear mixed model (GLMM)

- ▶ Regression model for clustered data where latent variables  $\eta_j$  induce within-cluster dependence of  $y_{ij}$  given covariates
  - Longitudinal: *i* is occasion or time-point (level 1), *j* is unit (level 2)
  - Cross-sectional: *i* is unit (level 1), *j* is cluster (level 2)
  - Say 'level-1 unit i' and 'level-2 unit j'
- ► Generalized linear mixed model (GLMM)

$$u_j = \mathbf{X}_j \boldsymbol{\beta} + \mathbf{Z}_j \boldsymbol{\eta}_j, \quad \boldsymbol{\eta}_j \sim N(\mathbf{0}, \boldsymbol{\Psi})$$

ightharpoonup Ex:  $n_i = 3$ ,  $\mathbf{Z}_i = \mathbf{X}_i \ (n_i \times 2)$ 

$$\underbrace{\begin{bmatrix} \nu_{1j} \\ \nu_{2j} \\ \nu_{3j} \end{bmatrix}}_{\mathbf{y}_{j}} = \underbrace{\begin{bmatrix} 1 & x_{1j} \\ 1 & x_{2j} \\ 1 & x_{3j} \end{bmatrix}}_{\mathbf{X}_{i}} \underbrace{\begin{bmatrix} \beta_{1} \\ \beta_{2} \end{bmatrix}}_{\boldsymbol{\beta}} + \underbrace{\begin{bmatrix} 1 & x_{1j} \\ 1 & x_{2j} \\ 1 & x_{3j} \end{bmatrix}}_{\mathbf{Z}_{i}} \underbrace{\begin{bmatrix} \eta_{1j} \\ \eta_{2j} \end{bmatrix}}_{\boldsymbol{\eta}_{j}}$$

GLLAMM – p.7

## Similarity of measurement model and GLMM

▶ Models for  $\nu_j$ 

Measurement model:  $u_j = \mathbb{I} \beta + \Lambda \eta_j$ 

Generalized linear mixed model:  $\nu_j = \mathbf{X}_j \boldsymbol{\beta} + \mathbf{Z}_j \boldsymbol{\eta}_j$ 

- ▶ Measurement model resembles GLMM where:
  - Items correspond to level-1 units and persons to level-2 units
  - Identity matrix  ${\mathbb I}$  replaces covariate matrix  ${\mathbf X}_j$
  - $\Lambda$  replaces  $\mathbf{Z}_i$ 
    - $\diamond$   $\Lambda$  does not vary between persons; elements are parameters
    - $\diamond$   $\mathbf{Z}_{j}$  varies between level-2 units; elements are fixed constants
- ▶ Will see that GLLAMMs unify measurement models and GLMMs

GLI AMM - p.9

# GLLAMM response model: Unifying latent variable models

 $\blacktriangleright$  Level-1 units *i* nested in level-2 units *j*, etc. up to level *L* 

$$\boldsymbol{\nu} = \mathbf{X}\boldsymbol{\beta} + \sum_{l=2}^{L} \sum_{m=1}^{M_l} \eta_m^{(l)} \mathbf{Z}_m^{(l)} \boldsymbol{\lambda}_m^{(l)}$$

Measurement model:

$$\underbrace{\begin{bmatrix} \nu_{1j} \\ \nu_{2j} \\ \nu_{3j} \end{bmatrix}}_{\pmb{\nu}_j} = \underbrace{\begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}}_{\pmb{X}_j} \underbrace{\begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix}}_{\pmb{\beta}} + \eta_{1j}^{(2)} \underbrace{\begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}}_{\pmb{Z}_{1j}^{(2)}} \underbrace{\begin{bmatrix} \lambda_1 \\ \lambda_2 \\ \lambda_3 \end{bmatrix}}_{\pmb{\lambda}_{1}^{(2)}} = \begin{bmatrix} \beta_1 + \eta_{1j}^{(2)} \lambda_1 \\ \beta_2 + \eta_{1j}^{(2)} \lambda_2 \\ \beta_3 + \eta_{1j}^{(2)} \lambda_3 \end{bmatrix}}_{\pmb{\beta}}$$

Generalized linear mixed model:

$$\underbrace{\begin{bmatrix} \nu_{1j} \\ \nu_{2j} \\ \nu_{3j} \end{bmatrix}}_{\boldsymbol{\nu}_{j}} = \underbrace{\begin{bmatrix} 1 & x_{1j} \\ 1 & x_{2j} \\ 1 & x_{3j} \end{bmatrix}}_{\boldsymbol{X}_{j}} \underbrace{\begin{bmatrix} \beta_{1} \\ \beta_{2} \\ \beta \end{bmatrix}}_{\boldsymbol{\beta}} + \eta_{1j}^{(2)} \underbrace{\begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix}}_{\boldsymbol{Z}_{1j}^{(2)}} + \eta_{2j}^{(2)} \underbrace{\begin{bmatrix} x_{1j} \\ x_{2j} \\ x_{3j} \end{bmatrix}}_{\boldsymbol{Z}_{2i}^{(2)}} = \begin{bmatrix} \beta_{1} + \eta_{1j}^{(2)} + (\beta_{2} + \eta_{2j}^{(2)})x_{1j} \\ \beta_{1} + \eta_{1j}^{(2)} + (\beta_{2} + \eta_{2j}^{(2)})x_{2j} \\ \beta_{1} + \eta_{1j}^{(2)} + (\beta_{2} + \eta_{2j}^{(2)})x_{3j} \end{bmatrix}}_{\boldsymbol{Z}_{2i}^{(2)}}$$

#### Exploratory latent class model

- ▶ Unit j falls into group (latent class) c with probability  $\pi_c$  (c = 1, ..., C)
- ► Each latent class characterized by different set of response probabilities for binary variables  $y_{ij}$  (i = 1, ..., n)
  - In medical diagnosis, latent classes are diseases and variables are diagnostic test results
- Exploratory latent class model
  - n-dimensional discrete latent variable  $\eta_j$  discrete values  $\{\mathbf{e}_1, \dots, \mathbf{e}_C\}$  with probabilities  $\{\pi_1, \dots, \pi_C\}$
  - Conditional response probabilities for class *c*:

$$\begin{aligned} \text{logit}[\text{Pr}(y_{ij} = 1 | \boldsymbol{\eta}_j = \mathbf{e}_c)] &= e_{ic} \\ \text{logit}[\text{Pr}(y_{ij} = 1 | \boldsymbol{\eta}_j)] &= \eta_{ij} \end{aligned}$$

GLMM with discrete random effects:

$$oldsymbol{
u}_j = \mathbf{I} oldsymbol{\eta}_j$$

## Response types in GLLAMM

Ordinary GLMs:

Links
identity
reciprocal
logarithm
logit
probit
compl. log-log

Families
Gaussian
gamma
Poisson
binomial

Extensions:

Ordinal responses
ordinal logit
ordinal probit
ordinal compl. log-log
scaled ordinal probit

Nominal & Rankings multinomial logit

- ▶ Mixed responses: different links and families for different units
- ► Heteroscedasticity: Variance or dispersion parameter  $\theta$  can be modelled as log  $\theta = \mathbf{z}^{(1)\prime} \alpha$

GLIAMM – p.11 GLIAMM – p.12

#### Structural model for continuous latent variables

► Regressions of latent variables on latent and observed variables at the same or higher levels

$$\eta = B\eta + \Gamma w + \zeta$$

$$\boldsymbol{\eta} = (\overbrace{\eta_1^{(2)}, \eta_2^{(2)}, \dots, \eta_{M_2}^{(2)}}^{\text{Level } l}, \underbrace{\eta_1^{(l)}, \eta_2^{(l)}, \dots, \eta_{M_l}^{(l)}}_{\text{Level } l, \dots, \eta_{M_l}^{(L)}, \dots, \eta_{M_l}^{(L)}, \dots, \eta_{M_L}^{(L)})'$$

- $M = \sum_{l} M_{l}$  latent variables:
- lackbox B is an upper triangular  $M \times M$  matrix of regression coefficients
- ightharpoonup is an  $M \times p$  matrix of regression coefficients
- ▶ w is a p dimensional vector of explanatory variables
- $\blacktriangleright$   $\zeta$  is an M dimensional vector of disturbances

$$\begin{split} \pmb{\zeta} &= (\overline{\zeta_1^{(2)}, \zeta_2^{(2)}, \dots, \zeta_{M_2}^{(2)}}, \dots, \overline{\zeta_1^{(l)}, \zeta_2^{(l)}, \dots, \zeta_{M_l}^{(l)}}, \dots, \overline{\zeta_1^{(L)}, \zeta_2^{(L)}, \dots, \zeta_{M_L}^{(L)}})' \\ \pmb{\zeta}^{(l)} &\sim N(\mathbf{0}, \pmb{\Psi}^{(l)}), \text{ independent across levels} \end{split}$$

GLLAMM - p.13

## Estimation and prediction in Stata program gllamm

- To obtain the likelihood of GLLAMM's, the latent variables must be integrated out
  - Sequentially integrate over latent variables, starting with the lowest level using a recursive algorithm
  - Use Gauss-Hermite quadrature to replace integrals by sums
  - Scale and translate quadrature locations to match the peak of the integrand using adaptive quadrature
- ▶ Maximum likelihood estimates obtained using Newton-Raphson
- ► Model-based and robust standard errors (sandwich estimator)
- ► Empirical Bayes (EB) predictions of latent variables and EB standard errors obtained as byproducts of adaptive quadrature

#### Structural model for discrete latent variables

▶ Probability that unit j belongs to class c = 1, ..., C may depend on covariates  $\mathbf{v}_i$  through a multinomial logit model

$$\pi_{jc} \; = \; rac{\exp(\mathbf{v}_j^\prime oldsymbol{arrho}^c)}{\sum_d \exp(\mathbf{v}_j^\prime oldsymbol{arrho}^d)},$$

where  $\varrho^c$  are parameters with  $\varrho^C = 0$  for identification

 Nonparametric maximum likelihood (NPML) estimator if number of classes increased to maximize likelihood

GLLAMM - p.14

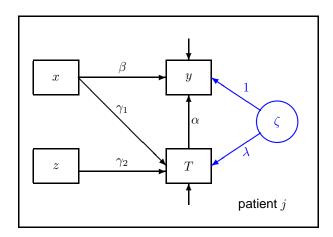
#### **Examples of GLLAMMs**

- Mixed response types
  - Ex 1: Random coefficient dependent dropout
  - Ex 2: Endogenous treatment model
- ▶ Continuous latent variables regressed on observed and latent var.
  - Ex 1: Item response model with differential item functioning (DIF)
  - Ex 2: Covariate measurement error model
  - Ex 3: Missing covariate model
- ▶ Discrete latent variables regressed on observed variables
  - Ex 1: Complier average causal effect (CACE) models
- Multilevel latent variables
  - Ex 1: Multilevel structural equation models (SEMs)

continue

GLIAMM – p.15 GLIAMM – p.16

# Endogenous treatment model

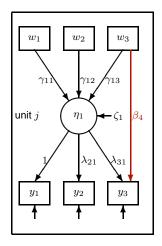


Treatment endogenous

back

GLLAMM - p.17

#### Item response model with DIF



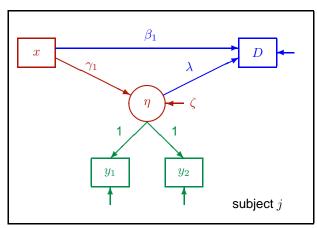
MIMIC model with direct effect

back

GLLAMM - p.18

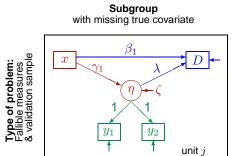
# Covariate measurement error

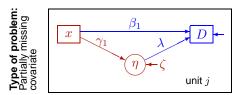
► Exposure model + Measurement model + Disease model



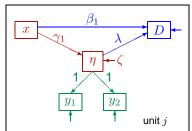
back

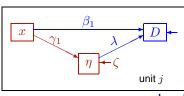
# Missing covariate model





Subgroup with complete data





back

GLLAMM - p.20

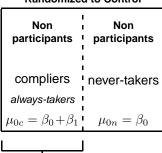
# Complier average causal model

# Randomized to Treatment

Non **Participants** participants compliers never-takers always-takers  $\mu_{1c} = \beta_0 + \beta_2$  $\mu_{1n} = \beta_0$ 

CACE

#### Randomized to Control



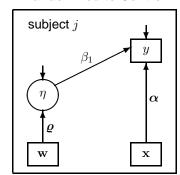
GLLAMM - p.21

# Complier average causal model (cont'd)

#### **Randomized to Treatment**

# subject j $\mathbf{x}$

#### Randomized to Control

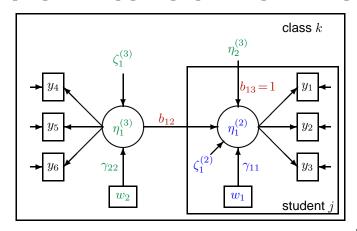


- ▶ c is dummy variable for complier (or always taker), versus never-taker
- $c=\eta$  is latent in control group (no access to treatment)
- Complier average treatment effect:  $\beta_2 \beta_1$

back GLLAMM - p.22

#### Multilevel SEM ► Structural model

$$\begin{vmatrix} \eta_{1jk}^{(2)} \\ \eta_{1jk}^{(3)} \\ \eta_{2k}^{(3)} \end{vmatrix} = \begin{bmatrix} 0 & b_{12} & 1 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} \eta_{1jk}^{(2)} \\ \eta_{1jk}^{(3)} \\ \eta_{2k}^{(3)} \end{bmatrix} + \begin{bmatrix} \gamma_{11} & 0 \\ 0 & \gamma_{22} \\ 0 & 0 \end{bmatrix} \begin{bmatrix} w_{1jk} \\ w_{2k} \end{bmatrix} + \begin{bmatrix} \zeta_{1jk}^{(2)} \\ \zeta_{1jk}^{(3)} \\ \zeta_{2k}^{(3)} \end{bmatrix}$$



# Multilevel SEM (cont'd)

Response model (measurement model)

$$\underbrace{\begin{bmatrix} \nu_{1jk} \\ \nu_{2jk} \\ \nu_{3jk} \\ \nu_{5k} \\ \nu_{6k} \end{bmatrix}}_{\boldsymbol{\nu}_{k}} = \underbrace{\mathbf{I}\boldsymbol{\beta}}_{\mathbf{X}_{k}} + \eta_{1jk}^{(2)} \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix}}_{\mathbf{Z}_{1k}^{(2)}} + \eta_{1k}^{(3)} \begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}}_{\mathbf{Z}_{1k}^{(3)}} \underbrace{\begin{bmatrix} 1 \\ \lambda_{2}^{(2)} \\ \lambda_{3}^{(2)} \\ \lambda_{1}^{(2)} \end{bmatrix}}_{\mathbf{Z}_{1k}^{(3)}} + \eta_{1k}^{(3)} \underbrace{\begin{bmatrix} 1 \\ \lambda_{2}^{(3)} \\ \lambda_{3}^{(3)} \\ \lambda_{3}^{(3)} \end{bmatrix}}_{\boldsymbol{\lambda}_{1}^{(3)}} + \eta_{2k}^{(3)} \underbrace{\begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\boldsymbol{\lambda}_{2}^{(3)}} \underbrace{\begin{bmatrix} 1 \\ \lambda_{2}^{(3)} \\ \lambda_{3}^{(3)} \end{bmatrix}}_{\boldsymbol{\lambda}_{1}^{(3)}} + \underbrace{\begin{bmatrix} 1 \\ \lambda_{2}^{(3)} \\ \lambda_{3}^{(3)} \end{bmatrix}}_{\boldsymbol{\lambda}_{2}^{(3)}} + \underbrace{\begin{bmatrix} 1 \\ \lambda_{2}^{(3)} \\ \lambda_{3}^{(3)} \end{bmatrix}}_{\boldsymbol{\lambda}_{3}^{(3)}} + \underbrace{\begin{bmatrix} 1 \\ \lambda_{2}^{(3)} \\ \lambda_{$$

▶ Column of zeros is necessary because  $\eta_{2k}^{(3)}$  has no measures

back

#### Concluding remarks

- ▶ Many more possible models in GLLAMM framework
- Advantage of modeling framework
  - Helps recognize commonalities between models
  - One program to estimate all models
  - Facilitates formulation and estimation of new models
- Features not covered in GLLAMM framework
  - Interactions among latent variables
  - Models with both discrete and continuous latent variables
- Problems with complex models
  - Difficult to establish identification and equivalence
  - Estimation can take a long time
  - Models may look naive but less naive than simpler (constrained) versions

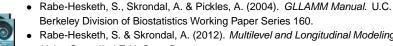
#### Some references related to GLLAMM

► Generalized Linear Latent and Mixed Modeling (GLLAMM) Framework:



GLI AMM - p.25

- Rabe-Hesketh, S., Skrondal, A. & Pickles, A. (2004). Generalized multilevel structural equation modelling. Psychometrika 69 (2), 167-190.
- Skrondal, A. & Rabe-Hesketh, S. (2004). Generalized latent variable modeling: Multilevel, longitudinal and structural equation models. Chapman & Hall/ CRC.
- Estimation and prediction:
  - Rabe-Hesketh, S., Skrondal, A. & Pickles, A. (2005). Maximum likelihood estimation of limited and discrete dependent variable models with nested random effects. Journal of Econometrics 128, 301-323.
  - Rabe-Hesketh, S. & Skrondal, A. (2006). Multilevel modelling of complex survey data. Journal of the Royal Statistical Society, Series A 169, 805-827.
  - Skrondal, A. & Rabe-Hesketh (2009). Prediction in multilevel generalized linear models. Journal of the Royal Statistical Society, Series A 172, 659-687.
- ▶ gllamm Software:



Berkeley Division of Biostatistics Working Paper Series 160.

• Rabe-Hesketh, S. & Skrondal, A. (2012). Multilevel and Longitudinal Modeling Using Stata (3rd Ed.). Stata Press.

► Check out http://www.gllamm.org for more information!

GLI AMM - p. 26